

DETERMINATION OF PARAMETERS AND EPIDEMIOLOGICAL FORECASTING USING MULTI-GENERATIONAL SIR MODELING

ABSTRACT. We use an age-dependent SIR system of equations to model the evolution of the COVID-19, taking different social and biological traits into account. Parameters that measure the amount of interaction in different locations (home, work, school, other) are approximated using a random optimization scheme, and indicate changes in social activities along the course of the pandemic. Taking previously determined parameters, we are able to make predictions for different scenarios. We compare our predictions with data from several locations in Brazil.

1. INTRODUCTION

Several months after the onset of the COVID-19 pandemic, it becomes clear how powerful numerical models can be used to predict different scenarios. In particular, compartmental epidemiological models are convenient, depending on few parameters and still capturing with reasonable precision essential aspects of the dynamics of infectious diseases.

The Susceptible-Infected-Recovered *SIR* modeling in its simplest form depends on two parameters, the average number of *adequate contacts* (those sufficient for transmission) and the *mean waiting period* in which the patient is infectious. It assumes that the population is homogeneous both in terms of behavior and biological susceptibility. It also assumes, in its simplest instance, that demographic aspects as age structure, natality and death rates do not matter for the dynamics of the disease.

Some of the above assumptions appear to hold well in the case of the COVID-19 pandemic. Demographic changes appear to be of less importance due to difference in time scales, compared with the evolution of the pandemic. Also, the period that one is infectious seems to be independent of age, to some reasonable extent. However, dropping the homogeneity hypothesis allows a more realistic model because age and location are certainly important in understanding the dynamics of the disease.

Mathematical epidemiology has a somewhat long history, with Hamer (1906); M'Kendrick (1925); Kermack et al. (1927) being among the earliest contributions. See Brauer (2017); Hethcote (2000) for a thorough review of mathematical modeling of infectious diseases,

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and Klepac and Caswell (2011); Klepac et al. (2009); Iannelli and Milner (2017); Li et al. (2020); Sun (2010); Towers and Feng (2012) for multi-group models. Some other models consider spatial aspects of infectious diseases, as Bertaglia and Pareschi (2020); Lang et al. (2018); Paeng and Lee (2017); Peixoto et al. (2020); Takács and Hadjimichael (2019); Viguerie et al. (2020). Theoretical aspects of models are covered in Kuniya and Wang (2018); Wen et al. (2018); Wu and Zou (2016). For the control related models for COVID-19 see for instance Perkins and España (2020) and references therein.

After the onset of the COVID-19 pandemic, a massive amount of articles related to the topic materialized. Several articles considered compartmental models to forecast different scenarios, as in Atkeson (2020); Prem et al. (2020); Komatsu and Menezes-Filho (2020); Viguerie et al. (2020); Walker et al. (2020).

Socio-economic costs of imposing social distancing are combined with compartmental models in Acemoglu et al. (2020); Alvarez et al. (2020); Atkeson (2020), and the interplay between economics and individual decisions are considered in Borelli and Góes (2020); Brotherhood et al. (2020); Eichenbaum et al. (2020). See Kremer (1996); d’Onofrio and Manfredi (2020); Funk et al. (2015); Manfredi and d’Onofrio (2013); Perrings et al. (2014); Soofi et al. (2020) for a description of behavioral- and economic- epidemiology ideas. Also, macro-economic aspects of the pandemic appears, e.g., in Guerrieri et al. (2020).

Population heterogeneity and its influence on herd immunity is considered in Aguas et al. (2020); Britton et al. (2020); Cui et al. (2019); Gomes et al. (2020). Vaccination is considered in, for instance, d’Onofrio and Manfredi (2020); González and Villena (2020); Jia et al. (2020); May and Anderson (1984).

Finally, alternative approaches, as agent-based and statistical models and a combination of these techniques with compartmental models (specially to gather data), are also commonly employed, as in Bendtsen Cano et al. (2020); Calvetti et al. (2020); Donnat and Holmes (2020); Ferguson et al. (2020); Roda et al. (2020).

In the present paper, we consider a modification of the age-structured SIR model used by Towers and Feng (2012). As in Prem et al. (2017), we allow for age-dependent probabilities of one being sub-clinical or clinical, and use their contact matrices which differentiates not only age but also location. Random searches approximate some key parameters that are related to social distancing in different locations, computing their values for some Brazilian regions. Following the approximation of those parameters, we are able to forecast best/worst case scenarios. To gauge the efficiency of the scheme, we compare with available data.

2. THE MODEL

In the spirit of compartmental modeling, we divide the whole population among *susceptible* (\mathbf{S}), *infected* (\mathbf{I}), *removed* (\mathbf{R}). Due to the short time scale of the disease, we can assume that demographic changes are not relevant, and deaths are estimated from the removed, \mathbf{R} . The above quantities are 16-dimensional vectors, stratified by age, running from 0-4 (first components), 5-9 (second components), etc, up to 75-above (16th components). There are P_i individuals for each age group, and the total population $P = \sum_{i=1}^{16} P_i$ is constant.

The equations governing the dynamics of the disease is as follows¹, for $i = 1 \dots, 16$ and $t \in \mathbb{N}$:

$$(1) \quad \begin{aligned} S_i(t+1) &= S_i(t) - \beta_i(\mathbf{I})S_i(t), \\ I_i(t+1) &= I_i(t) + \beta_i(\mathbf{I})S_i(t) - \gamma I_i(t), \\ R_i(t+1) &= R_i(t) + \gamma I_i(t), \end{aligned}$$

plus initial conditions. The model is “conservative” in the sense that $S_i + I_i + R_i = P_i$ is constant for all ages i .

The term $\beta_i S_i$ is the rate of new infections of the i -population, and we remark that β depends on the number of infected of all ages, making the model nonlinear. Those infected are removed from their condition at rate γI_i (they either recover or die). We further divide the infected into two groups, those that are asymptomatic or *sub-clinical* $I_i^{\text{sc}} = (1 - \rho_i)I_i$ in the sense that they do not require medical care. The *clinical* group $I_i^{\text{c}} = \rho_i I_i$ corresponds to those that get ill, requiring medical attention. The age-dependent fractions of those who become sub- or clinical are $\rho_i \in [0, 1]$.

To fully describe the model, we ought to define β and the other parameters.

(i) the components of the incidence function β are given by the formula

$$\beta_i(\mathbf{I}) = \sum_{j=1}^{16} \frac{\alpha^{\text{sc}} I_j^{\text{sc}} + \alpha^{\text{c}} I_j^{\text{c}}}{P_j} C_{i,j}^e = \sum_{j=1}^{16} (\alpha^{\text{sc}}(1 - \rho_j) + \alpha^{\text{c}} \rho_j) \frac{I_j}{P_j} C_{i,j}^e,$$

where C^e is an *effective contact matrix* that measures the number of contacts between age groups. It depends on the *amount* and *choices* of lock-down imposed, as we describe further ahead. The age-independent parameters α^{sc} and α^{c} correspond to the fraction of those infected that have the potential to infect others, due to social behavior or biology. For convenience, we also write the vector $\beta(\mathbf{I}) = C^e D \mathbf{I}$, where D is a diagonal

¹Note that we are *not* using the Einstein’s summation convention of summing up repeated indices.

matrix defined by

$$D_{jj} = \frac{\alpha^{\text{sc}}(1 - \rho_j) + \alpha^{\text{c}}\rho_j}{P_j}.$$

(ii) The time dependent effective contact matrix C^e is defined by

$$(2) \quad C^e(t) = \beta_h(t)C^{\text{home}} + \beta_w(t)C^{\text{work}} + \beta_s(t)C^{\text{school}} + \beta_o(t)C^{\text{other}},$$

where the exogenous data C^{home} , C^{work} , C^{school} , C^{other} are 16×16 matrices, where C_{ij}^ℓ indicates the average number of *different people* from the age group j that someone from the age group i contacted per day at $\ell \in \{\text{home, work, school, other}\}$, as compiled by Mossong et al. (2008); Prem et al. (2017). The parameters β_ℓ control the fraction of contacts that are “adequate”, in the sense that they lead to infections.

- (iii) Each component of $\boldsymbol{\rho}$ indicates whether an infected individual will become a sub-clinical or a clinical case. Since COVID-19 symptoms are more aggressive for older patients, ρ_i grows with i .
- (iv) $\gamma = 1 - e^{-1/d_I}$ indicates the daily probability that someone stays infected, where d_I is the average duration of infection, as in Prem et al. (2017).
- (v) The contagious parameters α^{sc} and α^{c} correspond to fractions of sub-clinical and clinical infected patients that still spread the virus.

In Acemoglu et al. (2020), a SIR model that takes into account age is also used, and three groups are considered. The authors include the possibility of using a family of scalings for the incidence function β , but none of them reproduces our own choice.

2.1. Reproduction and replacement numbers. The *basic reproduction number* \mathcal{R}_0 is understood as *the secondary cases produced by an infected individual introduced in a totally susceptible population*, see Barril et al. (2020); Delamater et al. (2019); Diekmann et al. (1990); Hethcote (2000). Hence, suppose that the whole population with distribution \mathbf{P} is susceptible ($\mathbf{S} = \mathbf{P}$), and we represent the infected individual by $\hat{\mathbf{I}}$, a vector belonging to the canonical basis $\{\mathbf{e}_1, \dots, \mathbf{e}_{16}\}$. The number of infected individuals in the first day belonging to the i -th is

$$\sum_{k=1}^{16} \delta_{ik} \beta_k(\hat{\mathbf{I}}) P_k = \sum_{k,j,\ell=1}^{16} \delta_{ik} P_k C_{kj} D_{j\ell} \hat{\mathbf{I}}_\ell = \sum_{m,j,\ell=1}^{16} \mathbb{P}_{im} C_{mj} D_{j\ell} \hat{\mathbf{I}}_\ell$$

where δ_{ik} indicates the Kronecker delta, which equals one if $i = j$ and zero otherwise, and we define \mathbb{P} as the diagonal matrix such that $\mathbb{P}_{kk} = P_k$. Let $A = \mathbb{P}CD$. Then the vector containing the contaminated individuals in the first day is $A\hat{\mathbf{I}}$, and we can bound its Euclidian

norm as

$$(3) \quad \|A\hat{\mathbf{I}}\|_2 \leq \lambda_{\max}\|\hat{\mathbf{I}}\|_2 = \lambda_{\max},$$

where λ_{\max} is the largest eigenvalue of A . Note that only the fraction $(1 - \gamma)^{d-1}\hat{\mathbf{I}}$ of the infected outsider remains infected on the d -th day. Those individuals infect less than $(1 - \gamma)^{d-1}\lambda_{\max}$, as in (3), and adding up all days, we gather that the total number of infected agents is bounded by the basic reproduction number

$$(4) \quad \mathcal{R}_0 := \sum_{d=1}^{\infty} (1 - \gamma)^{d-1} \lambda_{\max} = \frac{1}{\gamma} \lambda_{\max}.$$

Reasoning in a similar fashion, we define the *replacement number*

$$(5) \quad \mathcal{R}_t := \frac{1}{\gamma} \lambda_{\max}(t),$$

where $\lambda_{\max}(t)$ is defined as the maximum eigenvalue of the matrix $\mathbb{S}CD$ and \mathbb{S} is the diagonal matrix with the diagonal given by S_j for $j = 1, \dots, 16$.

The above definitions of \mathcal{R}_0 and \mathcal{R}_t rely on the Euclidian and spectral norms. However, if one is interested in considering age-dependent quarantine or vaccination policies, it might be interesting to consider the *infinity* norm. In this case, note that the total number of infected individuals on the first day is

$$(6) \quad \beta(\hat{\mathbf{I}}) \cdot \mathbf{P} = C^e D \hat{\mathbf{I}} \cdot \mathbf{P} = \sum_{i,j=1}^{16} C_{ij}^e D_{jj} I_j P_i \leq \|\mathbf{P}^T C D\|_{\infty},$$

where $\|\mathbf{v}\|_{\infty} = \max\{|v_i| : i = 1, \dots, 16\}$ for a vector \mathbf{v} of size 16. The above estimate holds since all quantities involved are non-negative. We remark that the estimate is sharp since, for at least one vector of the canonical basis, the inequality (6) becomes an identity. As before, adding up the days and considering that at the j th day there will be a fraction $(1 - \gamma)^{d-1}\hat{\mathbf{I}}$ of infected outsider, we gather that the total number of infected agents is bounded by

$$(7) \quad \mathcal{R}_{0,\infty} := \sum_{d=1}^{\infty} (1 - \gamma)^{d-1} \|\mathbf{P}^T C D\|_{\infty} = \frac{1}{\gamma} \max \left\{ \frac{(\alpha^{\text{sc}}(1 - \rho_j) + \alpha^c \rho_j)}{P_j} \sum_{i=1}^{16} C_{ij}^e P_i : 1 \leq j \leq 16 \right\}.$$

Similarly, we define the corresponding replacement number

$$(8) \quad \mathcal{R}_{t,\infty} := \frac{1}{\gamma} \max \left\{ \frac{(\alpha^{\text{sc}}(1 - \rho_j) + \alpha^c \rho_j)}{P_j} \sum_{i=1}^{16} C_{ij}^e S_i : 1 \leq j \leq 16 \right\}.$$

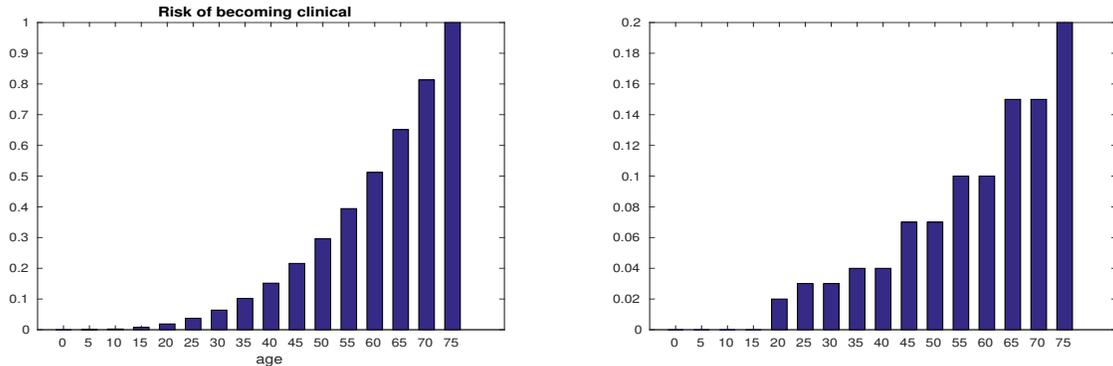


FIGURE 1. Probability of being clinical among infected patients ρ (left) and lethality weight w_d (right).

2.2. Modeling the lethality. Mortality is not part of the dynamics², resulting from the number of infected patients. We postulate that only those clinically infected die, at an age-dependent rate. So, among the “recovered” patients, $d(t) = \gamma\mu(t)\mathbf{w}_d \cdot \mathbf{I}^c(t)$ die, where \mathbf{w}_d is the lethality weight of the disease and the *letality strength* μ captures how letality varies with time. We assume that age-related lethality might vary across different regions and with time, but *the proportion* between different ages is constant.

2.3. Data. The transmission parameters α^{sc} and α^c are such that the transmission rate of asymptomatic is 60% larger than those symptomatic. This ad hoc choice is based on the assumption that symptomatic patients are more likely to quarantine themselves.

We stipulate that the proportion of symptomatic patients are as in Figure 1 (left), assuming that the proportion of symptomatic patients follows the same pattern of the infected that were hospitalized, as reported in Verity et al. (2020). However, the proportion of asymptomatic infected patients is far from being settled, with estimates ranging from 5% to 80%, as described in Heneghan et al. (2020).

Similarly, we assume that the lethality rates are as Figure 1 (right), and we assume that the proportions follow Brazeau et al. (2020); Verity et al. (2020).

The contact matrices C^{home} , C^{work} , C^{school} , C^{other} are from Prem et al. (2017), and population data, as seen in Figure 2 is from the Instituto Brasileiro de Geografia e Estatística (IBGE). The average number of days d_I that a patient is infectious is set to 12.

We use the available data (number of new cases and deaths) from each location to estimate the remaining functions $\beta_h(t)$, $\beta_w(t)$, $\beta_s(t)$, $\beta_o(t)$ used to define the effective matrix C^e in (2), and $\mu(t)$ defined in Section 2.2. These functions are transient since not only individuals

²We assume that the recovered patients do not get reinfected.

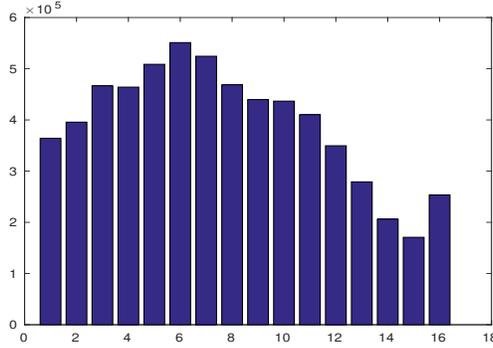


FIGURE 2. Population of the city of Rio de Janeiro.

change behavior with time, as in Brotherhood et al. (2020), but also lethality rates decay as the medical expertise increases with time.

For a simple SIR model, the available data uniquely determines the parameters, as in Bakhta et al. (2020). In general, that is not the case for our model, for instance if one of the four contact matrices vanish or if $C^{\text{work}} = C^{\text{other}}$. We make then some simplifying hypotheses. First, we assume that β_h is constant, as there is no reason to expect that the amount of contacts at home varies significantly with time. The second assumption is that $\beta_s(t)$ is identically zero, since schools were closed during the period of data gathering. We also suppose that $\beta_w = \beta_o$, i.e., the behaviour at work is equal to the behavior at other locations. Such assumption aims to reduce possible indeterminations of the parameters.

Consider that data is available for the period $\{1, \dots, T\} \subseteq \mathbb{N}$, and let $\mathcal{N}_{\mathcal{I}}^{\text{data}}(j)$ and $\mathcal{N}_{\mathcal{D}}^{\text{data}}(j)$ be the number of new cases and new deaths at the j -th day. For simplicity, assume that $T = N\delta t$ where $\delta t = 10$, and consider the partition $\mathcal{T}_{\delta t}$ of $\{1, \dots, T\} = \cup_{j=1}^N I_j$ where

$$\mathcal{T}_{\delta t} = \{I_j : j = 1, \dots, N\}, \quad I_j = \{(j-1)\delta t + 1, j\delta t\}.$$

Let $P_0(\mathcal{T}_{\delta t})$ be the space of piecewise constant functions with respect to the above partition.

To determine β_w we pose the problem of finding (β_h, β_w) that solves

$$(9) \quad \min_{(\hat{\beta}_h, \hat{\beta}_w) \in \mathbb{R} \times P_0(\mathcal{T}_{\delta t})} J(\hat{\beta}_h, \hat{\beta}_w), \quad \text{where} \quad J(\hat{\beta}_h, \hat{\beta}_w) = \frac{\|\mathcal{N}_{\mathcal{I}}^{\text{sir}}(\hat{\beta}_h, \hat{\beta}_w) - \mathcal{N}_{\mathcal{I}}^{\text{data}}\|_{\mathbb{R}^T}}{\|\mathcal{N}_{\mathcal{I}}^{\text{data}}\|_{\mathbb{R}^T}},$$

and $\|\cdot\|_{\mathbb{R}^T}$ is the Euclidian norm in \mathbb{R}^T . Above, $\mathcal{N}_{\mathcal{I}}^{\text{sir}}(\hat{\beta}_h, \hat{\beta}_w)(t)$ is the number of new infected clinical patients that the SIR model (1) yields at the t -th day if one replaces β_h and β_w by $\hat{\beta}_h$ and $\hat{\beta}_w$ in (2).

The above problem has finite dimension $N + 1$ since any function $\beta \in P_0(\mathcal{T}_{\delta t})$ can be written as

$$(10) \quad \beta_w(t) = \sum_{j=1}^N b_j \chi_{I_j}(t)$$

for some unknown coefficients $b_1, \dots, b_N \in \mathbb{R}$, where $\chi_{I_j}(t)$ is the characteristic function of I_j , which equals one if $t \in I_j$, and zero otherwise. The coefficients uniquely define a function in $P_0(\mathcal{T}_{\delta t})$ and vice-versa.

To solve (9) we employ a random optimization approach

- (i) Choose a initial guess (β_h, β_w)
- (ii) While some convergence criteria is not reached
 - (a) Define $(\hat{\beta}_h, \hat{\beta}_w)$ from (β_h, β_w) by adding noise
 - (b) If $J(\hat{\beta}_h, \hat{\beta}_w) < J(\beta_h, \beta_w)$, then $(\beta_h, \beta_w) \leftarrow (\hat{\beta}_h, \hat{\beta}_w)$

The estimation of the letality strength μ is based on the number of daily deaths data $\mathcal{N}_D^{\text{data}}$. The total number of deaths simulated by the SIR model is given by $\gamma\mu(t)\mathbf{w}_d \cdot \mathbf{I}^c(t)$. Equating both quantities we compute

$$(11) \quad \mu(t) = \frac{\mathcal{N}_D^{\text{data}}(t)}{\gamma\mathbf{w}_d \cdot \mathbf{I}^c(t)}.$$

3. SIMULATION OF THE DYNAMICS OF THE DISEASE

In what follows we show the results coming out of parameter estimations. We then show results addressing how well the model predicts the number of new cases. We compare with real daily data for new cases and deaths³. To smooth out oscillations, we present the data using a 7-day moving average.

One of the hurdles for models as ours is the definition of initial conditions, since when the virus started to spread and how many were the first infected individuals are not known. As a initial condition of our model, we assume that the first infected case occurred 30 days before the official recordings. Indeed, the first day of the official data logs 489 cases, and the real initial condition of the system is not known; cf. Lourenço et al. (2020), who speculates that the epidemic in Italy and UK started one month before the first reported death. We also estipulate that there were 10 infected patients at each age group at day one.

³All data were from taken from <https://covid.saude.gov.br/>, an official site from the Brazilian's Government Department of Health.

Such uncertainty of initial conditions makes the results unreliable at the initial periods of simulations. In particular, the computation of (11) is not feasible for the lack of data. We make then $\mu(j) = 1$ at the initial stages.

3.1. Parameter estimations, and prediction analysis. We consider results for the City of Rio de Janeiro, first presenting the computed results for all variables, adding up the values for all ages, i.e.

$$S(t) = \sum_{j=1}^{16} S_j(t), \quad I^{sc}(t) = \sum_{j=1}^{16} I_j^{sc}(t), \quad I^c(t) = \sum_{j=1}^{16} I_j^c(t), \quad R(t) = \sum_{j=1}^{16} R_j(t).$$

We recall the assumption that the contacts might lead to infections at home are time independent, and the contacts at work and other locations have a similar pattern, i.e., $\beta_w = \beta_o$. Also, schools are closed, thus $\beta_s = 0$.

We display in Figure 3 the results of our parameter estimations; see the caption for a detailed description of each individual plot. We note that the results from the model are accurate, approximating well the data and capturing the dynamics of the disease. In particular we see that β_w (top-right plot) oscillates following the “macroscopic” oscillations of daily cases (top-left plot). Regarding the letality, we note the decrease of its strength after day 90, in line with what was presented by Dennis et al. (2020); Horwitz et al. (2020).

Two important pieces of information that comes out of the modeling are the value of the reproduction and replacement numbers \mathcal{R}_0 and \mathcal{R}_t , and these computations follow easily from (4) and (5). We display the results in Figure 4 (left figure). Note that values of \mathcal{R}_0 is characterized only by the β 's and the data, while \mathcal{R}_t depends also on the dynamics of the disease through the susceptible population. That is important since even if $\mathcal{R}_0 > 1$, having \mathcal{R}_t smaller than one is enough to have a declining number of infected patients.

Another issue that is worth discussing is that we fit the number of clinically infected patients I^c to the data. However, we discussed nothing thus far about the sub-clinical patients I^{sc} . It turns out that the number of subclinical patients is, in this case, roughly ten times the number of registered patients, as shown in Figure 4 (right figure). Such difference resonates with often claims that there is a significant sub-notification. For instance, Havers et al. (2020) reports that the number of cases in the US might be at least 10 times of what is registered. It would not come as a surprise if this difference is even greater in Rio. Also, the WHO states that 80% of infections are mild or asymptomatic⁴, in particular among children,

⁴<https://www.who.int/emergencies/diseases/novel-coronavirus-2019/question-and-answers-hub/q-a-detail/q-a-similarities-and-differences-covid-19-and-influenza>

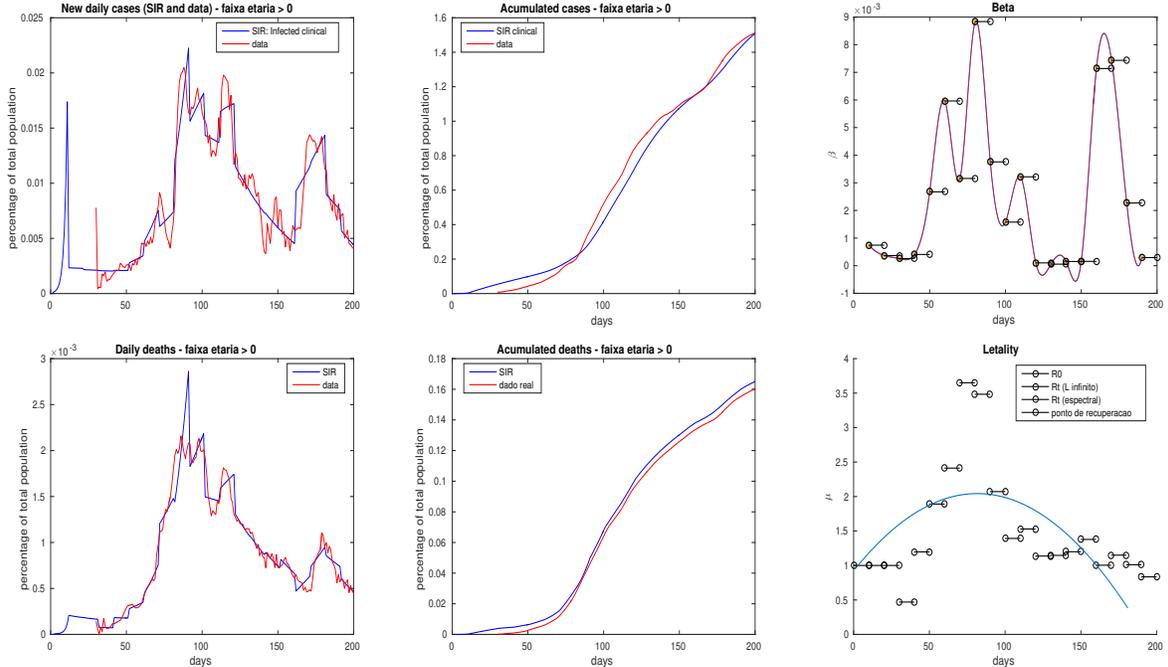


FIGURE 3. Results for the City of Rio de Janeiro. On the first couple of figures on the top we plot the new cases for each day, and the accumulated results. The red plot corresponds to real data, and the black plot depicts clinical cases modeled by SIR. On the top right we plot the estimated values of β_w , except for the first (that has a value roughly ten times bigger than the others). For visualization purposes we interpolated the discontinuous function β_w using cubic splines. The figures on the bottom are related to fatalities. The first and the second figures show daily and accumulated cases. The third figure displays the values of the letality strength μ , defined in (11). The blue curve is the least square quadratic curve.

as discussed by DeBiasi and Delaney (2020); Han et al. (2020). Recall that children are often not tested.

Next we investigate numerically if this model can make reasonable predictions. To do so, after the day 200, we run the simulations with frozen values of β_w and μ and compare with available data. We perform experiments considering the best and worst scenarios, i.e. lowest and highest number of cases and deaths. We do so by going back j 10-day periods and performing simulations with the highest and lowest values of β_w and μ in the period $200 - 10j$ and 200.

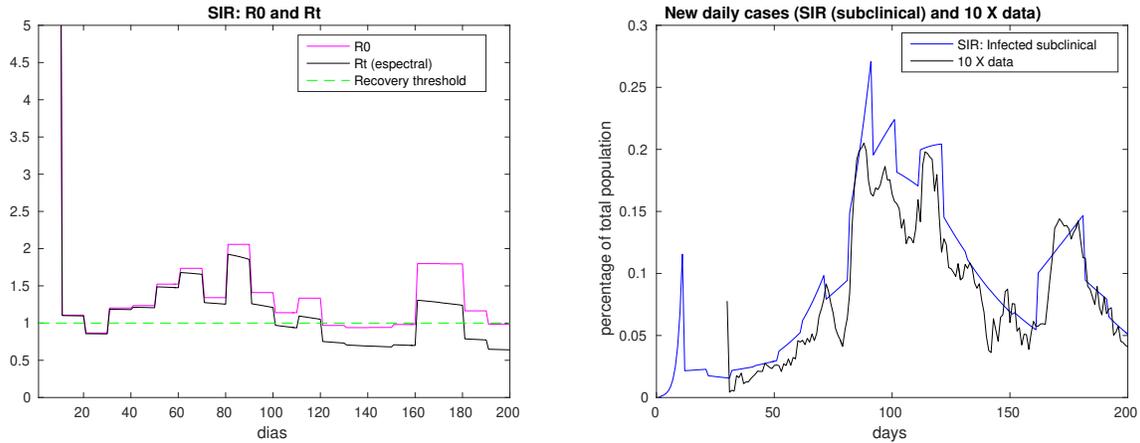


FIGURE 4. Reproduction and replacement numbers (left) and subclinical cases (right) for the city of Rio de Janeiro for the City of Rio de Janeiro.

Choosing $j = 4$, we show in Figure 5 the predicted number of new daily and acumulated cases and deaths with 50 days in advance. Note that, for most of the period of interest, the data is contained between the maximum and minimum value of predicted cases. We also plot results corresponding to the choice of β_w and μ being the average of the extreme values, yielding a more accurate prediction.

Regarding predictions for a long-term future (200 days), we plot in Figure 6 the acumulate cases for best, worst and average scenarios. We remark that the long term predictions in terms of number of deaths are pessimistic, since letality it is very likely that letality will continue to decrease in the future. And, of course, we do not consider the arrival of vaccines in our model.

3.2. Other locations. We next perform the same parameter estimation and prediction analysis for different regions. We show the respective results for the city of Petrópolis (Figure 7), State of Rio de Janeiro (Figure 8), City of São Paulo (Figure 9), State of São Paulo (Figure 10) and Brazil (Figure 11). Excepting for the number of cases for the whole country, the estimates are accurate enough. We remark that for Brazil’s case, the prediction are for 83 days in advance since more data is available.

4. DISCUSSIONS AND CONCLUSION

Epidemics as the COVID-19 are hard to model, not only because of biological factors, but also due to the unpredictability of human responses. Political, economical, social and individual factors influence the behavior of people, leading to various degrees of risk behavior.

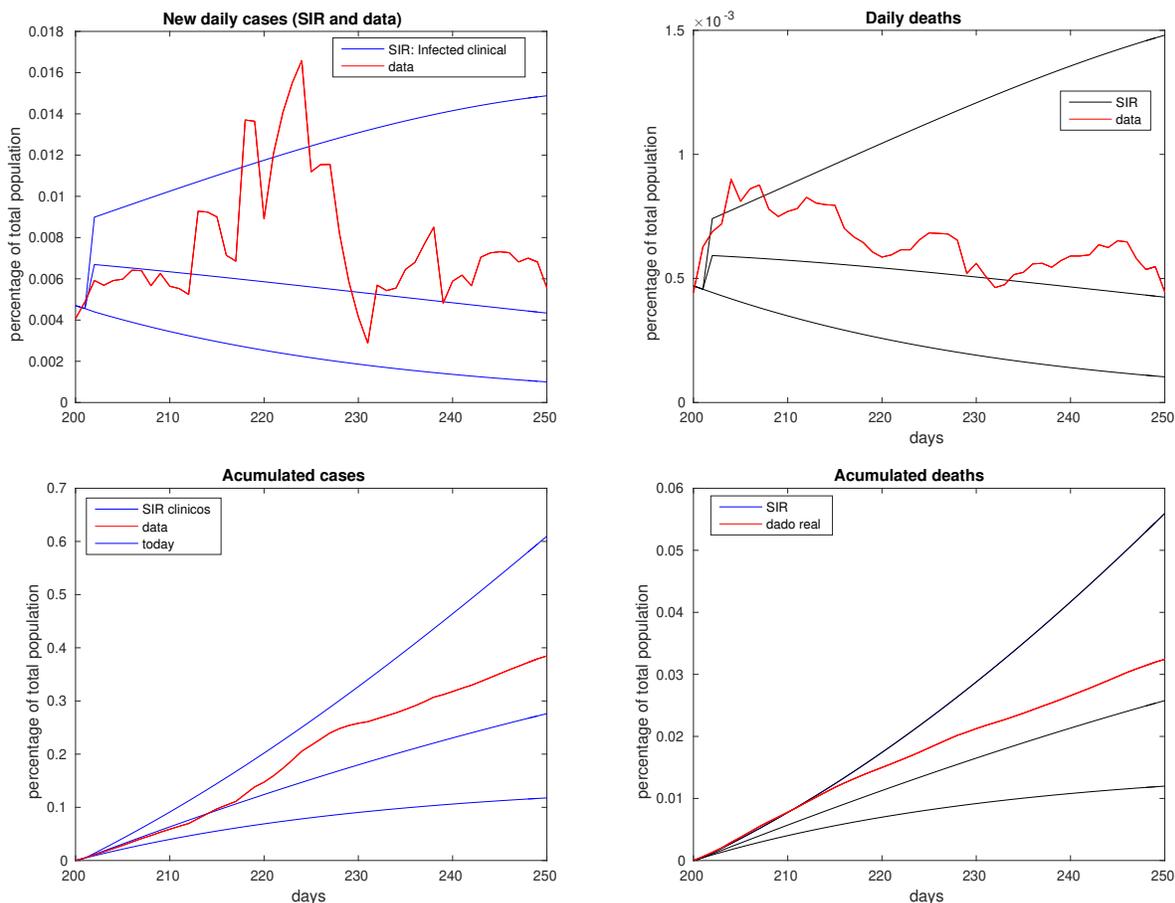


FIGURE 5. The top figures display the number of daily new cases and deaths, while the ones on the bottom display cumulative results. In all cases, the red plots indicates the data and blue represents the simulation results for frozen β_w and μ corresponding to the best and worst case scenarios.

Making long term predictions is riskier than usual in biological systems, but they are crucial for planning non-pharmaceutical interventions, allocation of resources, economical decisions, etc.

What we propose in the work is a general form to predict possible values of crucial parameters based on their past values. The main idea shares similarities with some Value at Risk analyses used in risk management, where the past dynamics of investments help in predicting the future.

Our method has two steps. Basing the dynamics of the disease on a multi-generational SIR model, we first determine the values of some time-dependent parameters, using a random optimization algorithm. Then, we feed the SIR model using extreme values of such

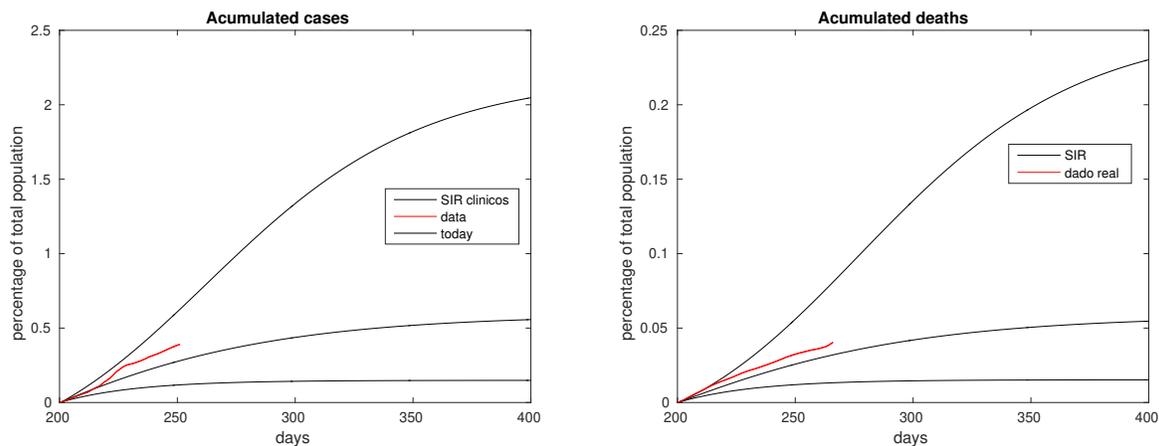


FIGURE 6. Long term predictions for accumulated cases (left) and deaths (right) in the City of Rio de Janeiro. As before, the data is on red, and the SIR model results are in black. In both figures, the top/bottom graphs display the worst/best case scenario. The middle graph displays the results when we take the average of the best/worst parameters.

parameters, and predict best/worst case scenarios. We are also consider a prediction that is between the two extreme cases.

We tested the method for two different states (Rio de Janeiro and São Paulo), and their capitals, predicting the dynamics for 50 days. In all tests, the method performed well, and the accumulated cases and deaths stayed within or very close to the bounds. We also tested for the whole Brazilian country for 83 days, getting good results for the accumulated deaths. For the accumulated cases, the results are not so good. That might be due to the extended period of prediction, or simply because the SIR model might not be appropriate for a large contry like Brazil.

Although we did not exploit this possibility, the SIR model used by us allows the investigation of “what-if” scenarios. For instance, what would happen if we open schools? Or what if we isolate the elderly population? What if we vaccinate the population of a predetermined age?

The methodology is general and can be applied to other systems and circumstances, and might be usefull to make long-term predictions in situations where unexpected occurances might change the behavior of the system.

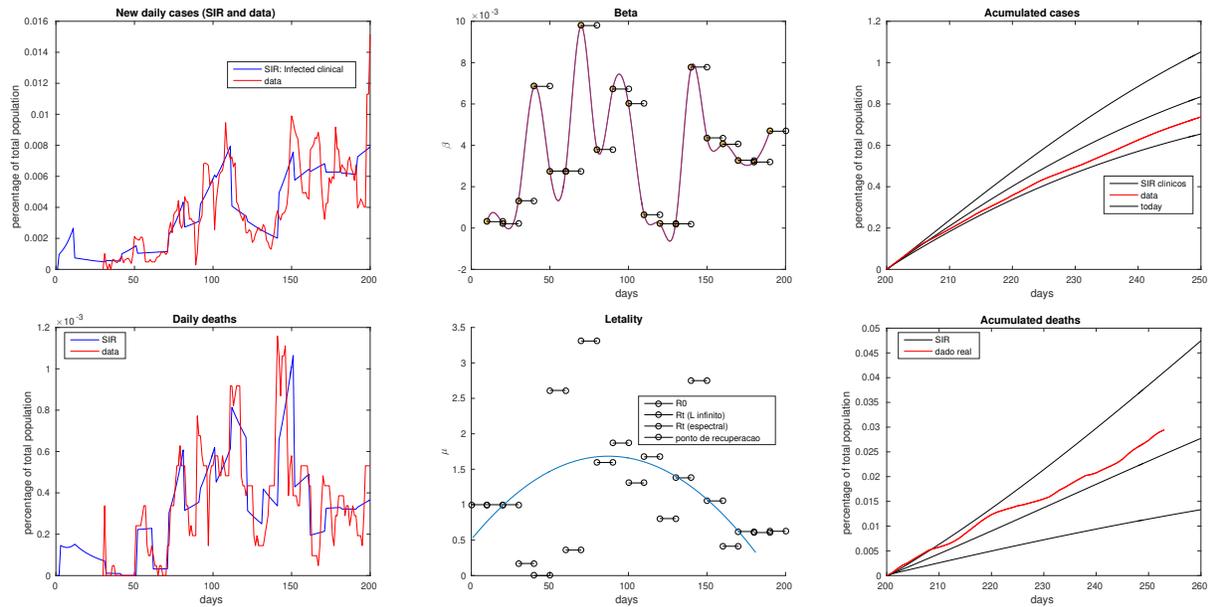


FIGURE 7. City of Petrópolis. Top left: new daily cases computed by SIR (blue) and data (red). Top middle: values of β_w . Top right: predictions for acumulated number of new cases; Bottom left: daily deaths computed by SIR (blue) and data (red). Bottom middle: μ (black) and blue curve is the least square quadratic parabola (blue). Bottom right: predictions for acumulated total deaths.

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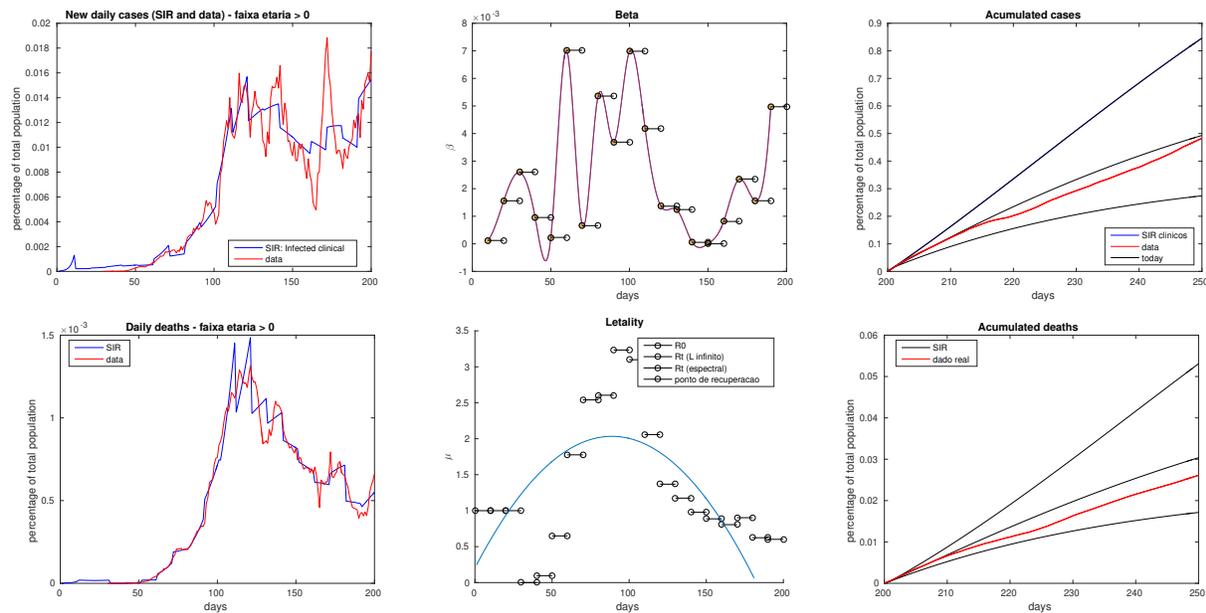


FIGURE 8. State of Rio de Janeiro. Top left: new daily cases computed by SIR (blue) and data (red). Top middle: values of β_w . Top right: predictions for acumulated number of new cases; Bottom left: daily deaths computed by SIR (blue) and data (red). Bottom middle: μ (black) and blue curve is the least square quadratic parabola (blue). Bottom right: predictions for acumulated total deaths.

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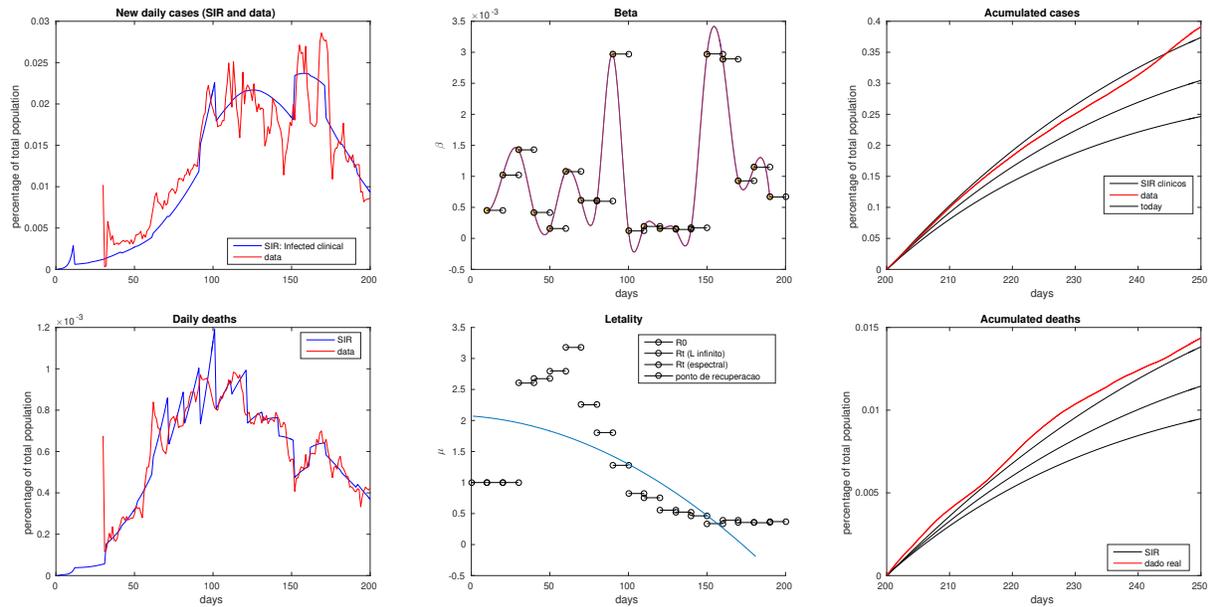


FIGURE 9. City of São Paulo. Top left: new daily cases computed by SIR (blue) and data (red). Top middle: values of β_w . Top right: predictions for acumulated number of new cases; Bottom left: daily deaths computed by SIR (blue) and data (red). Bottom middle: μ (black) and blue curve is the least square quadratic parabola (blue). Bottom right: predictions for acumulated total deaths.

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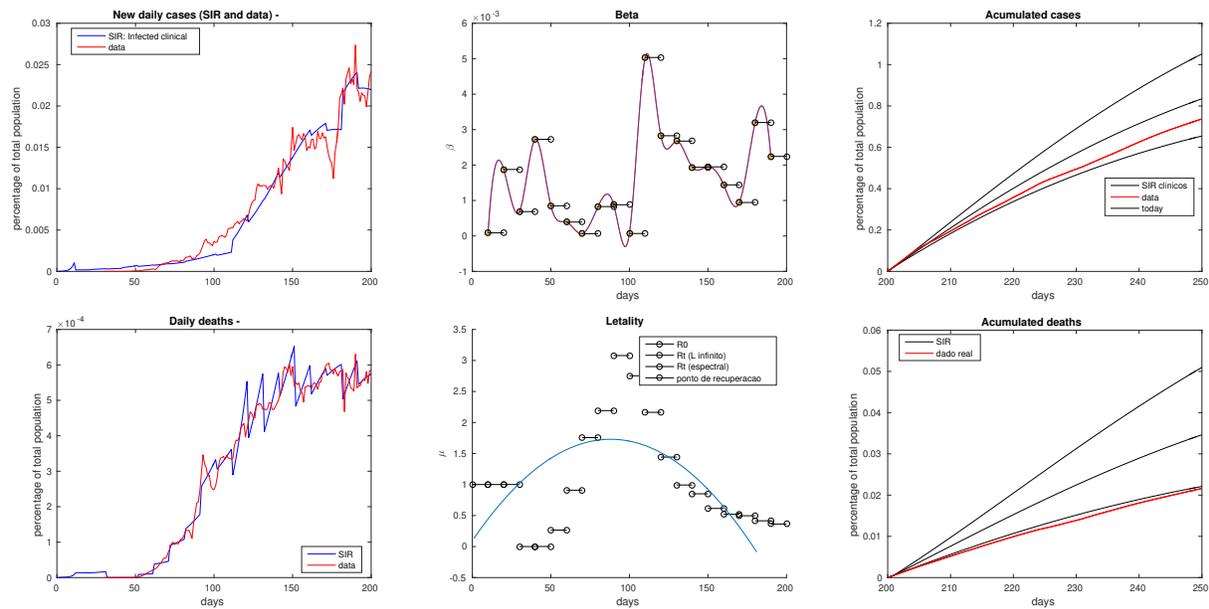


FIGURE 10. State of São Paulo. Top left: new daily cases computed by SIR (blue) and data (red). Top middle: values of β_w . Top right: predictions for acumulated number of new cases; Bottom left: daily deaths computed by SIR (blue) and data (red). Bottom middle: μ (black) and blue curve is the least square quadratic parabola (blue). Bottom right: predictions for acumulated total deaths.

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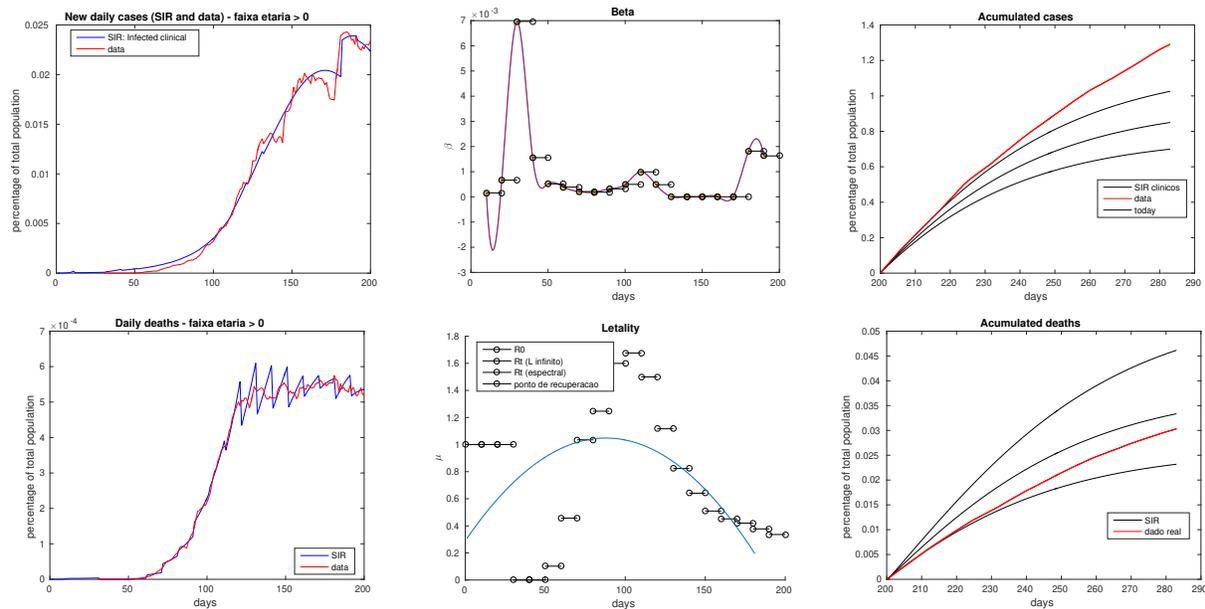


FIGURE 11. Brazil. Top left: new daily cases computed by SIR (blue) and data (red). Top middle: values of β_w . Top right: predictions for acumulated number of new cases; Bottom left: daily deaths computed by SIR (blue) and data (red). Bottom middle: μ (black) and blue curve is the least square quadratic parabola (blue). Bottom right: predictions for acumulated total deaths.

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